



Public Health Bulletin

Volume 7
Numbers 1-2
January-February, 1996

NSW  **HEALTH**
DEPARTMENT

ISSN 1034 7674
State Health Publication No (PH) 960024

SUICIDE MORTALITY IN NSW LOCAL GOVERNMENT AREAS

Gavin Stewart, Jennifer Chipps and Geoffrey Sayer
Mental Health Epidemiology,
Epidemiology Branch
Public Health Division, NSW Health Department

This is the fourth of a series of articles on the epidemiology of suicide in NSW by the Mental Health Epidemiology Group (MHEG). Our first report¹ presented Standardised Mortality Ratios (SMRs) for Area and District Health Services by pooling data across the 14-year period 1979-92. In this report we present the same analysis for Local Government Areas (LGAs). This is the finest level of geographic detail available in the Statewide death data.

We demonstrated in our first report that it is not feasible to use the suicide rate as a performance indicator for the success of suicide prevention programs in small populations, such as those of Area and (especially) District Health Services¹. This applies also to LGAs. However, it is common for questions to be asked about the suicide rate for a particular LGA, and the data here may be useful to those who are called on to answer questions about local suicide rates. We have grouped LGAs under the relevant Area or District Health Service and repeated the previous data for those Areas and Districts. This helps to indicate ways in which the data may be used to inform local discussions of suicide and prevention activity.

METHOD

Australian Bureau of Statistics suicide mortality data for Area and District Health Services and LGAs were pooled over the period 1979-92 and indirectly standardised relative to the pooled data for NSW during the same period. SMRs were scaled in the conventional way so SMR=100 for NSW as a whole. This standardisation adjusts for differences in the age-sex composition of the pooled population of the particular geographic area relative to the pooled NSW population. Thus an SMR of 120 means the particular area had a pooled suicide rate which was 20 per cent higher than that for NSW as a whole over the same period. Conventional 99 per cent confidence intervals were calculated for each SMR based on the usual assumption of Poisson-distributed events.

As some boundaries between LGAs have changed over time, we used aggregate data for LGAs where this was the case². Data for the LGAs of Central Sydney and Southern Sydney, Blacktown and Parramatta, Warringah and Pittwater, and a number of rural LGAs are presented for these groups, rather than for individual LGAs.

Most health administrative boundaries in NSW follow LGA boundaries. In general, therefore, it is possible to add up the observed and expected deaths for the set of LGAs reported under an Area/District heading, and arrive at the total for that Area or District. The exceptions to this are the composite LGA of Sydney and Southern Sydney (which divides in the proportion 40:60 between Central Sydney AHS and Eastern Sydney AHS³) and the LGA of Blayney (which is part of the composite LGA of Orange/Blayney/Cabonne because of boundary changes among these LGAs over time). We have reported data for Orange/Blayney/Cabonne under the Central Western DHS, but part of it belongs to the Evans DHS. Thus the data

Contents

Articles

1 Suicide mortality in NSW Local Government Areas

11 Waiting List Reduction Program, March to December 1995

Infectious Diseases

13 Notification trends

Correspondence

Please address all correspondence and potential contributions to:

The Editor,
NSW Public Health Bulletin,
Public Health Division,
NSW Health Department
Locked Bag No 961,
North Sydney NSW 2059
Telephone: (02) 391 9191
Facsimile: (02) 391 9029

Continued on page 2 ▶

Suicide mortality in NSW

► *Continued from page 1*

for each of the Areas and Districts are correct, and the data for individual LGAs and composites are correct, but the individual LGAs for Eastern and Central Sydney have to be combined to match the total deaths in those two areas. The same applies to Central Western and Evans DHS.

Recent changes in NSW Health administrative boundaries are only partly reflected in these data. We have reported the data for Canterbury LGA under both Southern Sydney AHS (where it forms part of the total for the AHS), and under Central Sydney AHS (where it now belongs, but does not, in this data, contribute to the AHS total). We have grouped Districts according to the Public Health Units which serve them, as in the previous report.

RESULTS

In each of Figures 1-6, the name of the Area or District or LGA is followed by the number of actual deaths by suicide over the 14-year period 1979-92, then the expected number of deaths for that locality if the NSW average age-specific death rates had applied to that population during the same period. The latter have been rounded to the nearest whole number.

Thus at the top of Figure 1 we show Central Sydney AHS with 737 actual deaths as against 570 which would have been expected for that population on State average rates. Immediately after the data for Central Sydney AHS are the data for the composite of Southern Sydney LGA and Central Sydney LGA: 328 actual deaths and 154 expected at State average rates. Of these, 40 per cent contribute to the totals for Central Sydney AHS. Data for LGAs are reported in descending order of SMRs within the Area or District.

For localities with at least five deaths, the graph shows the SMR and its 99 per cent confidence interval (CI). In the case of a few LGAs with very small populations, the value of the upper 99 per cent confidence limit exceeds SMR=300 and is not shown. The shaded area is the 'target' area of a 10 per cent reduction in the average suicide mortality for the State as a whole.

DISCUSSION

The data for 151 individual LGAs demonstrate, as previously, that estimates of the rates of rare events are very imprecise in small populations. There are various technical issues associated with the interpretation of these results, some of which were addressed in our first paper¹, and in correspondence arising out of it⁴. One of the agreed conclusions from that debate is that there is a need to offer guidance in the analysis and interpretation of health data of this kind.

For those reasons, we have added an account of some of the issues that arise when a large body of data is explored to detect "unusual" variation. The technicalities have been placed in footnotes as far as possible, and the main focus of the discussion is on the practical usage of the results.

Assessment of overall variation of LGAs relative to the NSW average

It is technically incorrect to compare indirectly standardised rates or SMRs with one another unless relatively sophisticated procedures are used^{5,6} and we have therefore interpreted each SMR in relation to SMR=100 for NSW as

a whole. As a practical decision rule, we used the 99 per cent CI surrounding the SMR estimate for an LGA to classify the LGA as "unusually" high or low relative to SMR=100⁷. Thus if SMR=100 lies below the lower 99 per cent CI bound for the SMR of the LGA, we consider that LGA as having an unusually high average suicide rate (for example, Marrickville, in Figure 1). Similarly, if SMR=100 lies above the upper 99 per cent CI bound for the SMR of the LGA, we consider that LGA as having an unusually low average suicide rate (for example, Hornsby, in Figure 1).

Using these definitions, there are eight LGAs with unusually high SMRs for suicide: Southern/Central Sydney, Leichhardt, Marrickville, Woollahra, Waverley, Randwick and North Sydney (Figure 1) and Broken Hill (Figure 4). There are also 10 LGAs with unusually low SMRs for suicide: Canterbury, Sutherland, Hornsby, Ku-ring-gai and Ryde (Figure 1), and Fairfield, Campbelltown, Baulkham Hills, Lake Macquarie and Port Stephens (Figure 2).

Clearly, there are other LGAs in Figures 1-6 which are "almost unusual", and not all the unusual LGAs are equally so. The confidence intervals shown in Figures 1-6 provide a much more informative view of these data than the simple classification based on an hypothesis-testing approach⁸. Nevertheless, in practice, confidence intervals tend to be used to classify groups of data, more or less as outlined above. It is therefore useful to estimate the probable errors involved in making decisions about geographical or administrative units in this way.

In classifying LGAs as unusually high or low, we concluded either that the result SMR=100 is within the range of likely values for the suicide rate in the LGA (negative decision – not "unusual"), or that it is not (positive decision – "unusual"). Because the test is based on a 99 per cent CI, we know there is a 1 per cent false positive rate for each individual decision. We made 151 individual decisions: one per LGA, and we can therefore expect that one or two "false positives" will result⁹. We can also calculate that the chance of obtaining a total of 18 "false positives" is essentially zero¹⁰. Taken together, these results mean we can safely conclude that there is some *systematic* variation in the standardised suicide mortality between LGAs in NSW, but we should not be preoccupied with the difference between the 18 "positives" (since one or two are likely to be *false* positives) and the remaining 133 LGAs.

Another useful way of looking at the data is to classify LGAs in relation to some meaningful SMR value other than SMR=100. We might, for example, accept as "less problematic" any LGA whose suicide rate can confidently be said to be no more than 20 per cent above the State average. These are the LGAs for which SMR=120 lies above the upper 99 per cent CI bound. This would add the LGAs of Botany, Hurstville, Kogarah, Warringah, Lane Cove (Figure 1), Bankstown, Camden, Wollondilly, Parramatta/Blacktown, Penrith, Hawkesbury, Newcastle, Wollongong, Shoalhaven (Figure 2), and Hastings (Figure 3) to the 10 LGAs with unusually low average suicide rates.

However, the most useful way of expressing the result for any particular LGA is in terms of its confidence interval. The most we can actually say is that, for example, the average suicide rate in Penrith LGA over the period 1979-92 was about 90 per cent of the State average, and there is a 99 per cent chance that the true value is between 75 per cent and 110 per cent of the State average. Data of this kind can be read off Figures 1-6 for every LGA in NSW.

Assessment of variation within health administrative areas

Most of the LGA data simply add detail to our previous conclusions about Area Health Services and Districts. The Central Sydney and Eastern Sydney Areas were classified as having unusually high SMRs in our previous analysis, and this is true of most (but not all) of the LGAs within them (Figure 1). The Far West District also has a high SMR, and this is clearly due to the LGA of Broken Hill (Figure 4). The Southern Sydney Area (Figure 1) and South Western Sydney Area (Figure 2) had unusually low SMRs in our previous analyses, and this is true of some of the LGAs within them.

The main value of the LGA-level analyses therefore lies in providing an indication to Area Health Services (and, to a lesser extent, District Health Services) of within-area variation. Looking at the variation in SMRs across LGAs in the same area gives an indication of whether the problem seems to be localised or uniform. For example, the Northern Sydney AHS as a whole is not different from the State average, but it contains an LGA (North Sydney) with one of the high rates and three LGAs (Hornsby, Ku-ring-gai, Ryde) with low rates.

The newly created South Eastern Sydney AHS will have an overall SMR closer to the State average than that of either of the original AHSs, because it will include four of the eight LGAs with high suicide rates (all from Eastern Sydney), and two of the 10 LGAs with low suicide rates (both from Southern Sydney). Clearly, the suicide problem is not solved by administrative mergers of this kind, but at the level of statistical summaries it might appear to be. For that reason alone it is useful to report data at a finer level of geographical detail. Similarly, the addition of Canterbury to Central Sydney AHS will reduce the suicide rate for the new AHS as a whole as the result of "immigration" of a lower-rate population, just as the "emigration" of that population from Southern Sydney AHS would have raised the average for the new area.

A practical way of looking at this is to consider the estimated SMRs for the LGAs within an Area or District against the 99 per cent CI for the Area or District as a whole. This violates the rule about not comparing indirectly standardised SMRs with one another⁵, but is a reasonably valid way of assessing the uniformity of the suicide rate across geographical subdivisions of a larger area. For example, in the Northern Sydney AHS (Figure 1), there are four LGAs (Hunters Hill, Manly, North Sydney, Mosman) whose SMRs lie above the upper 99 per cent confidence limit for the AHS as a whole, and another four (Hornsby, Lane Cove, Ku-ring-gai, Ryde) whose SMRs lie below the lower limit. On this basis, those concerned with suicide prevention in the Northern Sydney AHS might choose to focus on the former group of LGAs. By contrast, the data for the Wentworth, Central Coast and Illawarra Area Health Services (Figure 2) show quite uniform SMRs across the relevant LGAs, and an Area-wide approach might be more suitable.

It may also be noted that the LGAs with the highest SMRs in the Northern Sydney AHS would be regarded as "average" in Central Sydney AHS and are below the average in Eastern Sydney AHS, but issues of this kind are better dealt with by considering the Area or District Health Service as a whole in relation to the NSW average.

Measures to set and monitor progress towards targets

A number of useful measures may also be derived from the data to assist local discussions on suicide issues, and in planning how to set and monitor progress towards targets. These measures are of most value to District Health Services where populations are small, annual data are very variable and sophisticated statistical advice may be difficult to obtain. We will illustrate some uses of the reported data by referring to the Murrumbidgee DHS (Figure 5).

Next to each locality in Figures 1-6 we have presented the number of actual deaths, and the number of deaths expected if the NSW average age-specific suicide rates applied through the years 1979-92. The ratio of these (actual/expected), expressed as a percentage, is the SMR. Thus for the Murrumbidgee DHS the SMR can be calculated¹¹ as (SMR=actual deaths as percentage of expected deaths = 79/76 x 100 = 104). The plotted point scaled off Figure 5 is about 105.

It is also useful to present the same evidence as "excess mortality". For the Murrumbidgee DHS as a whole there would have been 76 deaths by suicide in 14 years if people in that District were at the same risk as others in the State as a whole. In fact there were 79 deaths, so the excess mortality of the District due to suicide amounted to three deaths in 14 years. This was not uniform across the District, however, since the individual LGA data show there were four excess deaths in Hay, five in Leeton and one in Narrandera, while Griffith had six fewer deaths than expected, Carrathool one fewer and Murrumbidgee had the same suicide mortality as the rest of the State. To avoid differences of this kind being over-emphasised, it should be noted that "excess mortality" is the difference between actual and expected deaths, while an SMR is the ratio of actual to expected deaths, expressed as a percentage, which is equal to 100 when actual deaths equals expected deaths. Therefore, since the SMRs for all the individual LGAs in the Murrumbidgee DHS, and for the DHS as a whole, have wide confidence intervals which include SMR=100, it follows that the corresponding confidence intervals for the excess mortality are wide, and include excess mortality=zero¹².

Similar calculations can be used to determine the expected number of deaths for other SMR values. Suppose the District administration chose to aim at a reduction to SMR=90, as indicated in the target range. What actual number of deaths would correspond to SMR=90? This can be determined in the following way. SMR=100 corresponds to the expected number of deaths (76). Therefore SMR=90 corresponds to 90 per cent of the expected number of deaths (90/100 x 76 = 68.4). This is the number of expected deaths over 14 years. This corresponds to a change from 79/14 = 5.6 per year to 68.4/14 = 4.9 per year. For the Murrumbidgee DHS, therefore, a target of SMR=90 relative to the NSW average rate between 1979-92 corresponds to 5.6 - 4.9 = 0.7 deaths a year. This target (one life saved a year) is easier to relate to than achieving a reduction in the SMR from 105 to 90. It also may help to consider the relative benefit of investments in this area of health care as against other life-saving interventions.

Calculations of this kind can readily be carried through for any SMR that is chosen as a target, for any of the data in Figures 1-6. It must be remembered, though, that these are 14-year data for the period 1979-92, so even though the data

Continued on page 4 ►

Suicide mortality in NSW

► *Continued from page 3*

are more reliable, they are not very sensitive to recent population shifts or changes.

CONCLUSIONS

The main value of the LGA-level analyses lies in providing indications to those concerned with suicide prevention in Area and District Health Services. This may help in explaining the issues at the local level, in targeting programs on smaller administrative areas which have higher SMRs than others, and perhaps in developing programs with local government. In addition, LGA-level data are useful in avoiding the possibility that health problems which are distributed unevenly in the population can become invisible when various sub-populations are merged and only aggregate data are reported. This is well illustrated by the recent merging of the former Eastern Sydney and Southern Sydney Area Health Services (less Canterbury LGA) to create the South-Eastern Sydney AHS. There is, therefore, a clear role for Public Health Units in addressing these issues of within-area variation in health status indicators.

There is also a value in understanding the statistical basis of interpreting observed data. The suicide risk for a population results from a multiplicity of factors – some chronic, some acute, some affecting the population as a whole, and some affecting only a few individuals within it. In a fairly typical LGA in NSW, with a population of 20,000 males, constant over the period 1979-92, there were four male deaths a year on average, but in some years as many as eight and in one year no deaths at all. The only "pattern" in the data which would allow the number in any particular year to be predicted is that it was consistent with a Poisson distribution with a mean of four deaths a year. If we take this as a model, we know there is only about one chance in 100 of 10 or more deaths occurring if the underlying risk has remained constant. Between the unobservable risk, and the observable data, lie all the unknown factors which affect the individuals in a population. The power of a statistical model is that it makes efficient use of all the available information to guide decisions and to provide indicators against which we can measure suicide prevention programs. In particular, it gives us an estimate of the extent to which the observed data will vary even when the underlying risk remains the same.

By contrast, if we do not take this into account when we assess the performance of suicide prevention programs against observable indicators, the decisions we make on programs may be as random as the year-to-year variation in suicide deaths in a typical LGA in NSW. The alternative is not to abandon the attempt to address these problems on a local level, but to choose appropriate indicators of performance. Programs which have a logical relationship to risk reduction in the population, and which are agreed by relevant expert opinion – including the opinion of community groups – to be of value, need to be implemented and sustained for long enough to have an effect which is detectable. Every reasonable effort must be made to establish a connection between program activity and its impact on risk, even if it is impossible to carry the argument through to a demonstrable impact on the suicide rate.

ACKNOWLEDGMENTS

We wish to thank Dr Tim Churches, Epidemiology Branch, NSW Health Department, for supplying the SAS routine which was adapted for the standardisation, and Kerrie Crawford and Margaret Riakos from the Centre for Mental Health for verifying the data in the figures.

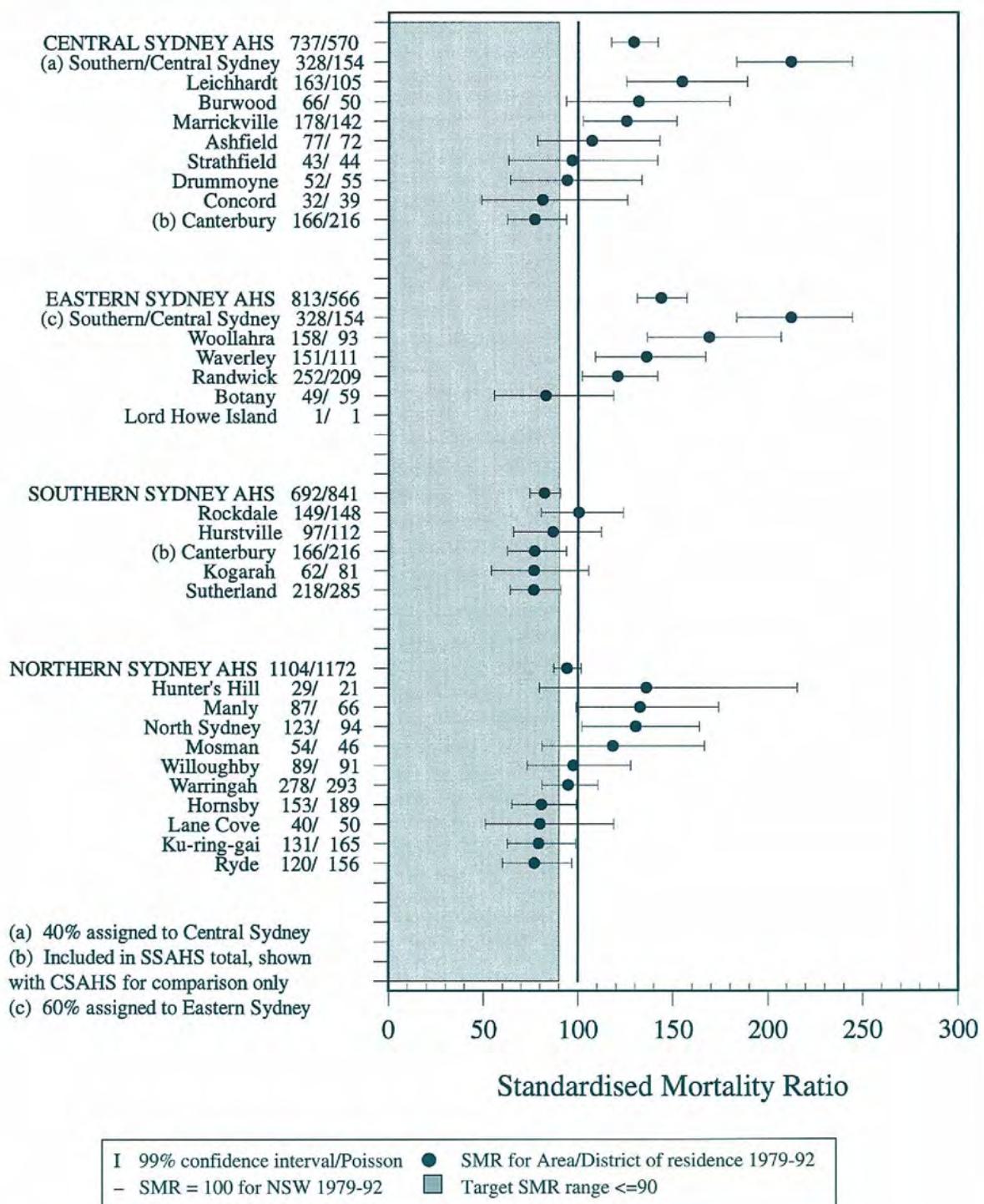
NOTE

Membership of MHEG is open to people with a professional interest and expertise in mental health epidemiology who are willing to contribute to the planning and production of a series of publications and reports on important mental health topics. The policy of MHEG is joint publication by the group as a whole in which authors are listed in order of their contribution to the particular report. The contact address for matters concerning this report is Mental Health Epidemiology Group, Centre for Clinical Policy & Practice, Public Health Division, NSW Health Department, Locked Bag 961 PO North Sydney 2059 (Fax: 391-9232, Internet e-mail gstew@gwsm.doh.health.nsw.gov.au).

1. Stewart G, Chipps JA, Sayer G. Suicide mortality in NSW: geographic variations. *NSW Public Health Bulletin*, 1995; 6(6):49-52.
2. The mortality and population data used in this report contain codes for "continuous" LGAs. These are either individual LGAs whose boundaries have not changed over the period 1970-1995, or combinations of LGAs in cases where the inter-LGA boundary has changed, but the boundary for the aggregate has not. These data were prepared for the HOIST (Health Outcomes Information and Statistical Toolkit) system by Dr Tim Churches and his group, based on similar files developed by the NSW Central Cancer Registry.
3. This apportionment has been done by a particular randomised allocation of deaths in the mortality file in the HOIST data system, since the place of residence of the person at death is only classified to the LGA. On average the assignment is in the ratio 40:60, but in any year or group of years it will not be exactly 40:60 because of the randomisation. This problem arises only because of the arbitrary boundary chosen to divide these LGAs for NSW Health administrative purposes.
4. The discussion can be found in the *NSW Public Health Bulletin*, 1995; 6(9):92-94. The essential point is that our confidence intervals may, in some circumstances, be underestimates of the true width of the 99 per cent confidence intervals.
5. Carrieri KC and Roos L. Comparing standardised rates of events. *Am J Epidemiol* 1994; 14:472-482.
6. Chan LY, Gibberd RW. Statistical properties of an index to measure systematic regional variation in mortality. Unpublished paper, October 15, 1991.
7. Technically, we are not classifying the result for the LGA as "high" or "low", but rather concluding either that the result SMR=100 was within the range of likely values for the suicide rate in the LGA, or it was not.
8. Gardner MJ, Altman DG (eds). *Statistics with confidence*. London: *British Medical Journal*, 1992.
9. If there is a 1 per cent chance of a false positive for each test, and 151 tests, the expected number of false positives is 1.51, so "one or two" is a practical summary of this result. There is also a false negative rate, which would need to be based on some chosen level of SMR judged "important to detect". For example, if it were considered important to detect SMR=200 (that is, that the suicide rate in a geographical area was twice the State average), it would be possible to estimate the false negative rate for each area. This would be different for each, and depend on the population, but clearly the false negative rates will be high because the confidence intervals at the LGA level are so wide. Even when the true rate was equivalent to SMR=200, many of the confidence intervals would still include SMR=100.
10. If the chance of a single false positive is 0.01 (1 per cent), then the binomial probability of obtaining more than 12 false positives in a series of 151 tests is less than 10⁻⁶.
11. For those who may wish to report the SMR at the highest accuracy available, greater precision will be obtained by scaling off the point plotted in the relevant figure for small numbers of deaths (because the figure given for expected deaths has been rounded to the nearest whole number) and will be obtained by calculating from the actual and expected death data otherwise.
12. If we cannot conclude that the SMR for the District differs from SMR=100, then neither can we conclude that excess mortality differs from zero.

FIGURE 1

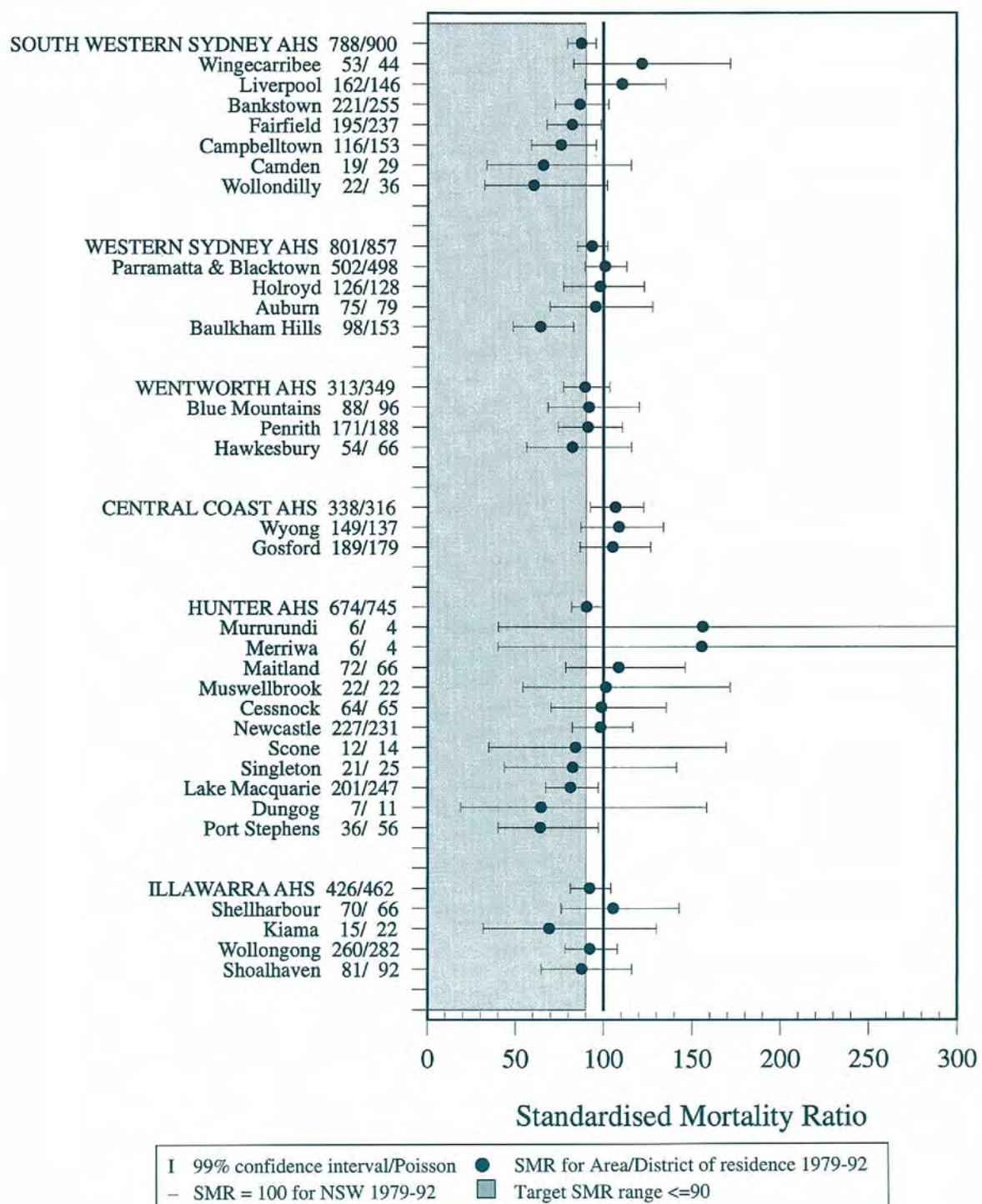
LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92



Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

FIGURE 2

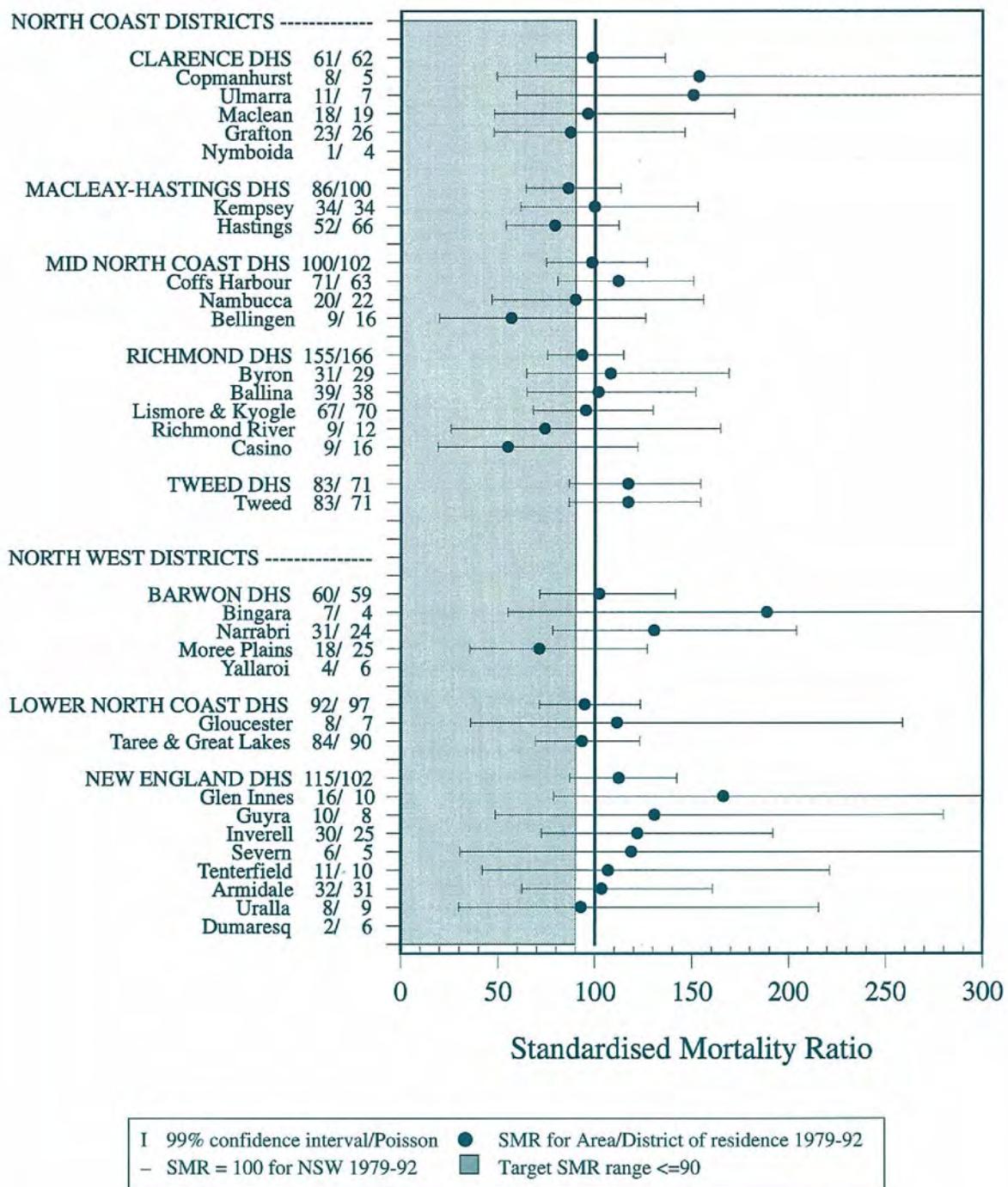
LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92



Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

FIGURE 3

LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92



Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

FIGURE 4

LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92

NORTH WEST DISTRICTS (cont'd)-----

NORTH WEST DHS 108/113

Walcha	8/ 6
Barraba	6/ 5
Tamworth	53/ 50
Manilla	5/ 5
Quirindi	8/ 9
Parry	12/ 17
Gunnedah	14/ 20
Nundle	2/ 2

WESTERN DISTRICTS -----

CASTLEREAGH DHS 52/ 41

Mudgee	30/ 23
Coonabarabran	14/ 11
Coolah	8/ 7

FAR WEST DHS 95/ 48

Broken Hill	86/ 41
Central Darling	7/ 5
Unincorp Far West	2/ 2

MACQUARIE DHS 120/105

Cobar	15/ 8
Warren	9/ 6
Dubbo	55/ 45
Wellington	17/ 14
Bogan	6/ 6
Coonamble	7/ 9
Gulgandra	6/ 8
Narromine	5/ 10

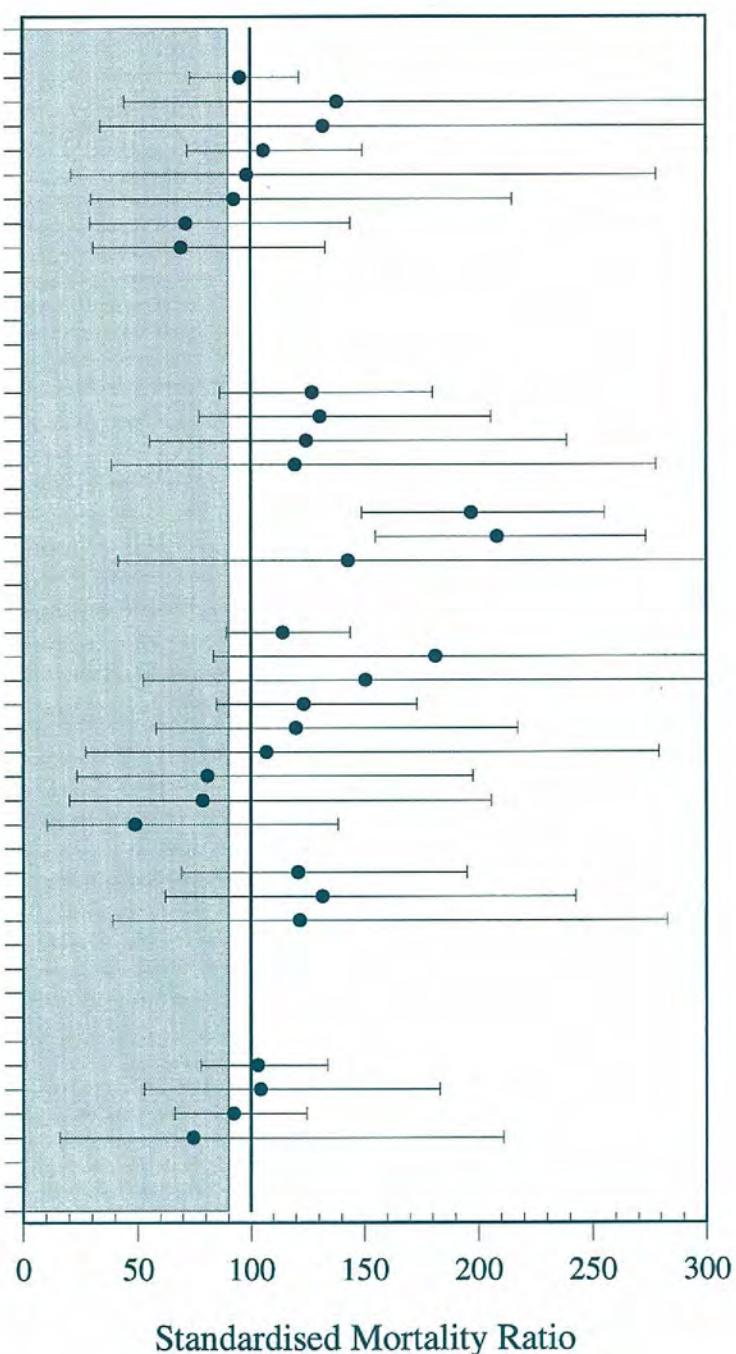
ORANA DHS 27/ 22

Walgett	16/ 12
Bourke	8/ 7
Brewarrina	3/ 4

CENTRAL WEST DISTRICTS -----

CENTRAL WESTERN DHS 94/ 91

Cowra	19/ 18
Orange & Blayney & Cabonne	70/ 76
Weddin	5/ 7



I 99% confidence interval/Poisson
- SMR = 100 for NSW 1979-92
■ SMR for Area/District of residence 1979-92
■ Target SMR range <=90

Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

FIGURE 5

LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92

CENTRAL WEST DISTRICTS (cont'd)

EVANS DHS	85/ 99
Rylstone	10/ 7
Bathurst & Evans & Oberon	49/ 51
Greater Lithgow	26/ 32

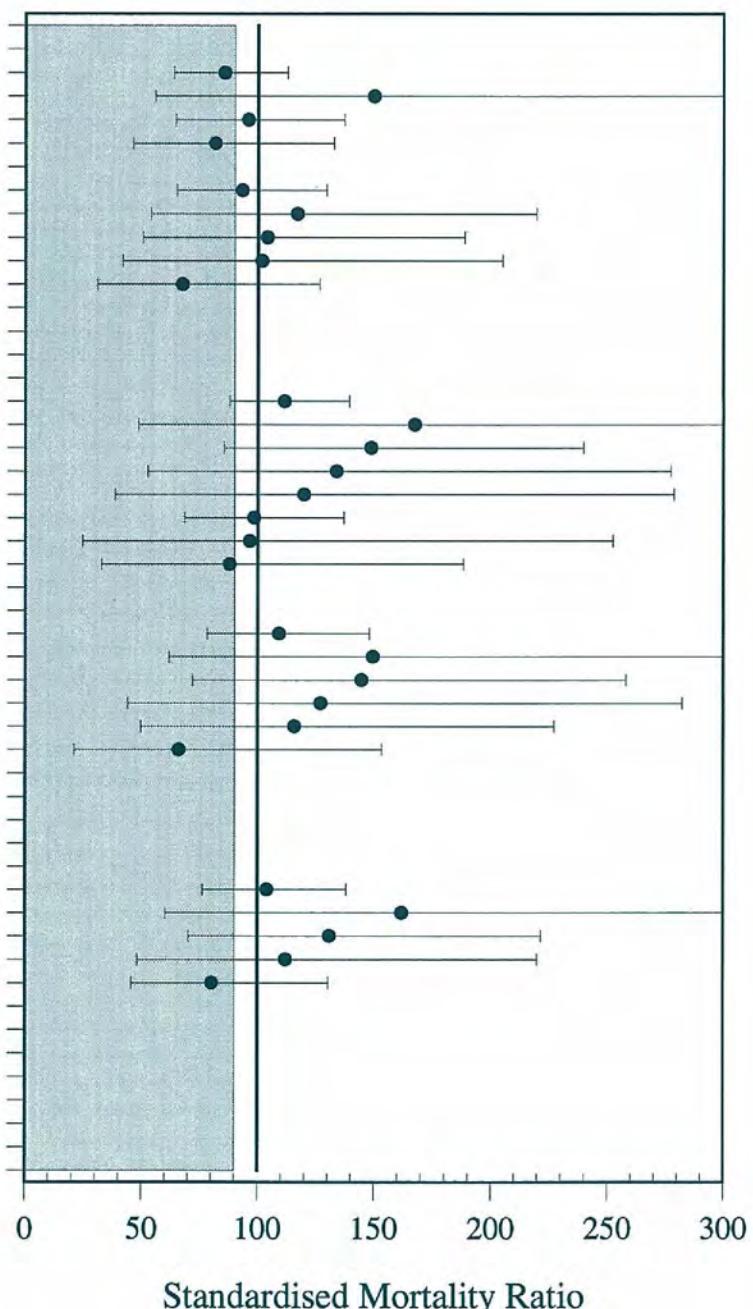
LACHLAN DHS	59/ 63
Lachlan	15/ 13
Forbes	17/ 16
Bland	12/ 12
Parkes	15/ 22

SOUTH WEST DISTRICTS -----

HUME DHS	130/116
Holbrook	7/ 4
Tumut	27/ 18
Hume	11/ 8
Culcairn	8/ 7
Albury	58/ 59
Tumbarumba	6/ 6
Corowa	10/ 11
Urana	3/ 3

MURRAY DHS	68/ 62
Wakool	12/ 8
Deniliquin	18/ 12
Murray	9/ 7
Wentworth	13/ 11
Berrigan	8/ 12
Balranald	4/ 4
Conargo	1/ 3
Jerilderie	3/ 3
Windouran	0/ 1

MURRUMBIDGEE DHS	79/ 76
Hay	10/ 6
Leeton	22/ 17
Narrandera	13/ 12
Griffith	26/ 32
Carrathool	4/ 5
Murrumbidgee	4/ 4

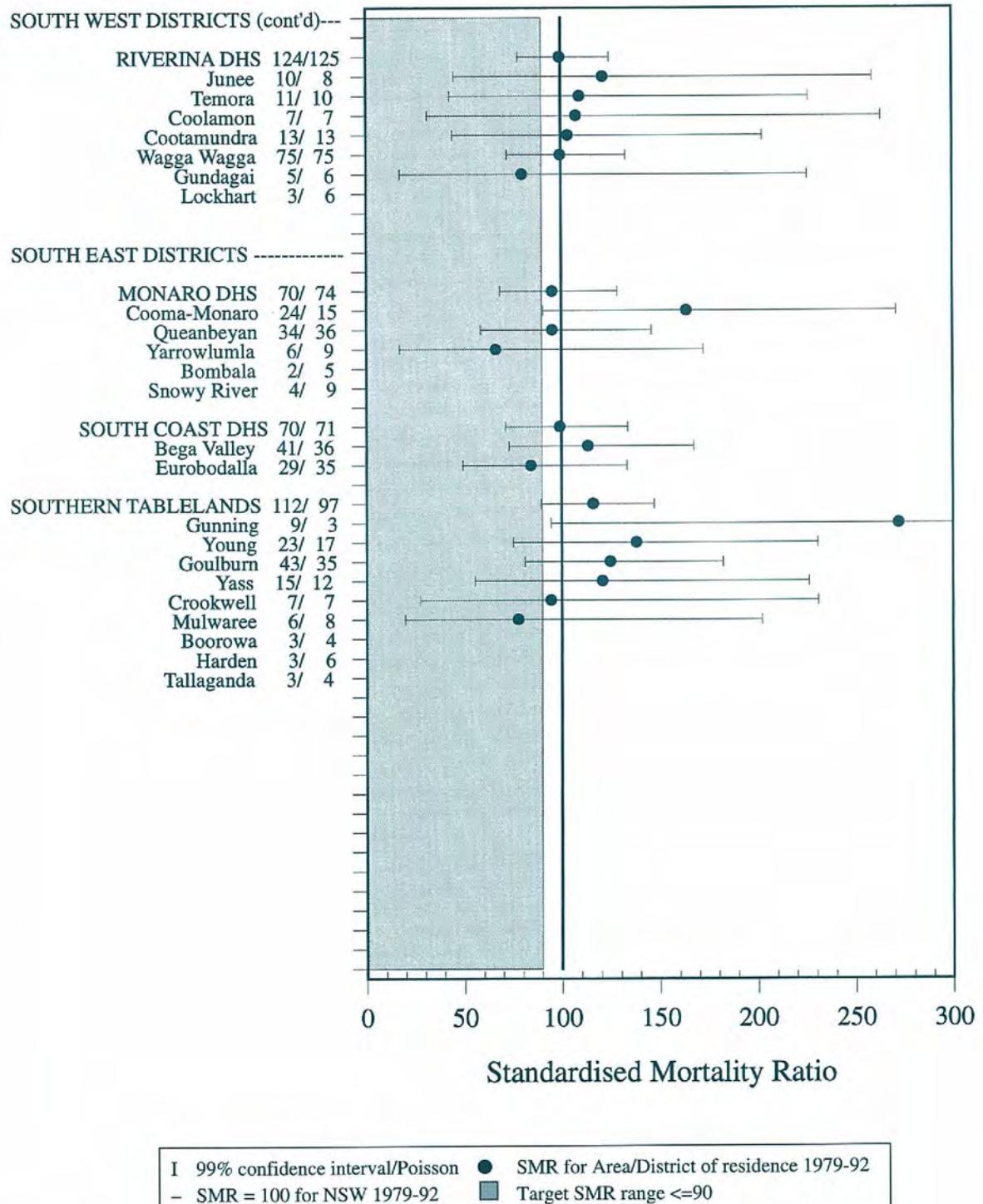


I	99% confidence interval/Poisson	●	SMR for Area/District of residence 1979-92
-	SMR = 100 for NSW 1979-92	■	Target SMR range <=90

Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

FIGURE 6

LOCALITY, OBSERVED DEATHS/NSW AVERAGE AND SMRS FOR SUICIDE PERSONS NSW 1979-92



Source: Epidemiology Branch, ABS deaths 1979-92, registered by 1993

WAITING LIST REDUCTION PROGRAM

MARCH TO DECEMBER 1995

Nick Shiraev and Jean McGarry
Evaluation and Monitoring Branch
Performance Management Division
NSW Health Department

The Waiting List Reduction Program was introduced as an initiative to improve access to elective surgery services in public hospitals throughout NSW. The main aim was to halve elective surgery waiting lists within 12 months. Particular emphasis was to be placed on those people who had been waiting more than six months for surgery.

The Minister for Health held meetings with a range of health professionals who indicated their support for the program. A State task force was formed, including representatives of the health professions, to oversee and advise on the program. The role of the task force was to monitor and guide progress towards achieving the principal targets of the program, in particular the establishment of best practice principles. It also dealt with clinical and workforce issues.

The success of the Waiting List Reduction Program would not have been possible without the partnership developed between NSW Health, the health professions and unions.

At the inception of the program, each Area and District Health Service developed an implementation plan. Local areas of greatest need were identified and priorities set. Specialists, general practitioners, nurses, allied health professionals, support staff and managers were involved in consultation at hospital, Area, District and State levels, and they formed Area and District steering committees.

There was extensive liaison with Areas and Districts over the development of their plans, with support provided by the NSW Health Department.

The program was an ambitious undertaking involving enormous commitment from health professionals. As far as is known, no program of this kind has been successfully undertaken anywhere in Australia or overseas.

RESULTS

Between March 31 and December 31, 1995 the elective surgery waiting list in NSW decreased by 25,118 patients, or 56.2 per cent – down to 19,589 – thus attaining the 50 per cent reduction target of the Waiting List Reduction Program launched in May 1995 (Table 1). All Areas and Districts, except four, exceeded their 50 per cent target (Table 2).

The combined medical and surgical list was reduced by 50.0 per cent – from 52,740 in March (44,707 surgical and 8,033 medical) to 26,365 (19,589 surgical and 6,776 medical) at the end of December.

By December, all Areas and Districts except four reduced their surgical lists by more than 50 per cent. Areas and the larger Districts with relatively big waiting list reductions were Western Sydney (68.5 per cent), Northern Sydney (62.9 per cent), South Eastern Sydney (59.8 per cent), Hume (76.3 per cent), New England (73.3 per cent), North West (60.9 per cent) and Central Western, despite shortage of operating theatres until September, (59.0 per cent). The Greater West (Western, South Western Sydney and Wentworth Areas) reduced its lists by 6,605 – or 58.5 per cent.

- The number of patients waiting more than 12 months fell by 78 per cent from 2,265 to 497. The number of patients waiting between six and 12 months fell by 71 per cent, from 6,379 to 1,850. This amounted to a reduction of 6,297 – or 73 per cent – (down from 8,644 to 2,347) of all patients waiting longer than six months.
- Expected waiting time shortened by half a month to 40 days.
- The average time on list declined by one month to 2.5 months. The average waiting time for those admitted during December dropped by nine days to 36 days.

Waiting lists, expected waiting times and average waiting times decreased in each specialty.

There were major reductions in the number of people waiting more than six months for the following key procedures: cataract extraction (80 per cent), coronary artery bypass graft (95 per cent), myringotomy (80 per cent), tonsillectomy (75 per cent), total hip replacement (50 per cent), total knee replacement (35 per cent) and varicose vein stripping and ligation (90 per cent).

Elective surgery admissions increased by 26,019 – or 14.4 per cent – over the corresponding period last year.

NEW EMPHASIS

The Waiting List Reduction Program to date has achieved a substantial reduction in the number of patients on surgical waiting lists and it is now appropriate to be directing the focus of attention towards waiting times. The NSW Health Department, in consultation with the NSW Branch of the Royal Australasian College of Surgeons and Area and District Health Services, is developing benchmarks for the length of time patients should wait for their surgery.

Benchmarks will reflect the priorities of the new waiting list focus. They will cover the wait of both urgent and long-wait patients.

Benchmarks will also be developed with an aim of reducing disruptions to patients which occur as a result of rescheduling admissions following unforeseen demand. The emergency workload of a hospital is difficult to predict. When this is under-estimated, elective patients may need to be rescheduled. The aim is to minimise the number of delays through improved scheduling systems. In this area, performance indicators would cover the number of delays relative to the number of admissions.

The program has resulted in longer-term improvements such as increases in peri-operative and early discharge programs. These programs enable improvements in service, such as patients being admitted on the day of their operation.

Other benchmarks are being proposed to improve patient management. These include:

- development of benchmarks for same day surgery;
- reduction of unplanned returns to theatre;

Continued on page 13 ▶

TABLE 1**ELECTIVE SURGERY WAITING LISTS
MARCH TO DECEMBER, 1995**

Month	Number on list (excl. list transfers)	% Reduction	Number waiting 6-12 months	% Reduction	Number waiting >12 months	% Reduction
March	44,707		6,379		2,265	
June	42,163	6	5,689	11	2,074	8
July	39,628	11	4,785	25	2,059	9
August	37,341	17	4,235	34	1,783	21
September	34,299	23	3,762	41	1,428	37
October	29,780	33	2,705	58	922	59
November	24,701	45	2,132	67	637	72
December	19,589	56	1,850	71	497	78

TABLE 2**REDUCTIONS IN WAITING LISTS BY AREA AND DISTRICT HEALTH SERVICE,
MARCH TO DECEMBER, 1995**

Elective Surgery Waiting List Reduction Program, NSW Health Department

AREA/DISTRICT	NUMBER WAITING				TOTAL REDUCTION	
	Mar 31, 95 Original list	Dec 31, 95 Including transfers*	Transfers* April-December	Dec 31, 95 Excluding transfers*	March to December	
					Number	%
Central Sydney	2,779	1,668	327	1,341	-1,438	-51.7
Northern Sydney	2,708	1,150	144	1,006	-1,702	-62.9
Western Sydney	4,650	2,165	698	1,467	-3,183	-68.5
Wentworth	2,122	1,075	18	1,057	-1,065	-50.2
South West Sydney	4,514	2,392	235	2,157	-2,357	-52.2
Central Coast	2,317	1,646	516	1,130	-1,187	-51.2
Hunter	4,178	2,547	101	2,446	-1,732	-41.5
Illawarra	2,778	1,324	-	1,324	-1,454	-52.3
South Eastern Sydney	7,190	3,288	395	2,893	-4,297	-59.8
Royal Alexandra Hospital	792	485	66	419	-373	-47.1
Barwon	43	3	-	3	-40	-93.0
Castlereagh	51	9	9	-	-51	-100.0
Central Western	588	321	80	241	-347	-59.0
Clarence	199	15	-	15	-184	-92.5
Evans	330	92	19	73	-257	-77.9
Far West	204	88	-	88	-116	-56.9
Hume	413	160	62	98	-315	-76.3
Lachlan	111	48	-	48	-63	-56.8
Lower North Coast	1,071	555	52	503	-568	-53.0
Macleay-Hastings	385	138	40	98	-287	-74.5
Macquarie	813	484	47	437	-376	-46.2
Mid North Coast	887	422	28	394	-493	-55.6
Monaro	59	27	-	27	-32	-54.2
Murray	16	-	-	-	-16	-100.0
Murrumbidgee	-	105	105	-	-	0.0
New England	404	108	-	108	-296	-73.3
North West	1,213	589	115	474	-739	-60.9
Orana	33	1	-	1	-32	-97.0
Richmond	681	395	71	324	-357	-52.4
Riverina	1,276	620	-	620	-656	-51.4
South Coast	89	172	172	-	-89	-100.0
Southern Tablelands	187	66	43	23	-164	-87.7
Tweed Valley	824	456	95	361	-463	-56.2
Port Macquarie Base Hospital	802	413	-	413	-389	-48.5
New South Wales	44,707	23,027	3,438	19,589	-25,118	-56.2

Source: DOHRS, January 11, 1996

* List transfers, in general, are an administrative change and not a change in local demand, and are therefore not taken into account when estimating the CHANGES in the number of patients on a list.

INFECTIOUS DISEASES

Waiting List Reduction Program

► Continued from page 11

- development of benchmarks for theatre utilisation; and
- reduction of unplanned readmissions.

The new emphasis on waiting times will not only involve inpatient care, but also outpatients and community-based health services. Although much is known about waiting times for inpatient treatment, information on how long people wait for services in outpatient clinics or community health services is limited. It is clear, however, that some people are waiting unacceptably long periods for services such as podiatry, physiotherapy and speech therapy. The Department is undertaking a survey to provide information in this area.

AUDITS

Between March and December 1995, Area and District Health Service compliance with Departmental policy on waiting lists was continually audited by the State Waiting List Auditor.

In addition, an audit and review of NSW public hospital waiting lists conducted by Coopers and Lybrand found, *inter alia*:

“... no evidence of any material errors or direct manipulation of data ... The figures, results and qualitative statements included in the report were arrived at in accordance with the methodology outlined in the survey and are consistent with previous reports” and “... have been recorded in accordance with the Health Department policies and procedures on waiting lists ...”

“The source data used for reporting purposes was consistent with information contained in DOHRS and WLCOS (Departmental Information Systems). This data was also consistent with unit record information maintained at Area/District Health Service level.”

PUBLIC HEALTH EDITORIAL STAFF

The editor of the Public Health Bulletin is Dr Michael Frommer, Director, Research and Development, NSW Health Department. Dr Lynne Madden is production manager.

The *Bulletin* aims to provide its readers with population health data and information to motivate effective public health action. Articles, news and comments should be 1,000 words or less in length and include a summary of the key points to be made in the first paragraph. References should be set out using the Vancouver style, the full text of which can be found in *British Medical Journal* 1988; 296:401-5.

Please submit items in hard copy and on diskette, preferably using WordPerfect, to the editor, NSW Public Health Bulletin, Locked Mail Bag 961, North Sydney 2059. Facsimile (02) 391 9029.

Please contact your local Public Health Unit to obtain copies of the *NSW Public Health Bulletin*.

This issue of the *NSW Public Health Bulletin* contains information on infectious disease notifications for December 1995 and January 1996, as well as cumulative notifications for the 1995 year received to January 31, 1996.

NOTIFICATION TRENDS

In December 1995 notification levels were higher than historical levels for hepatitis A and rubella (Figure 7). Notification trends for rubella were discussed in the August, September and October issues of the *NSW Public Health Bulletin*.

Notification rates were lower than historical levels in December 1995 for foodborne illness (not otherwise specified), gastroenteritis, *Haemophilus influenzae* type b (Hib) infection, measles, meningococcal disease, pertussis and Q fever (Figure 7).

HEPATITIS A

There has been a steady increase in hepatitis A notifications from the inner Sydney area since October 1995 (Figure 8). Of the 162 notifications for hepatitis A received for the period December 1995-January 1996, 82 were residents of Eastern Sydney. Central Sydney Public Health Unit reported a similar but less marked increase.

COMMITTEES OF THE AIDS/INFECTIOUS DISEASES BRANCH

The AIDS/Infectious Diseases Branch provides the secretariat for several key advisory committees on communicable disease issues. The committees bring together a wide range of expertise and their deliberations guide the Department in the development of policy and legislative change. This section of the *Bulletin* will report from time to time on their recommendations. The committees include:

CAS	Ministerial Advisory Committee on AIDS Strategy
HAC	Hepatitis Advisory Committee
ICAG	Infection Control Advisory Group
IDAC	Infectious Diseases Advisory Committee
IAC	Immunisation Advisory Committee
SHAC	Sexual Health Advisory Committee
TBAC	Tuberculosis Advisory Committee
LSAC	Laboratory Surveillance Advisory Committee

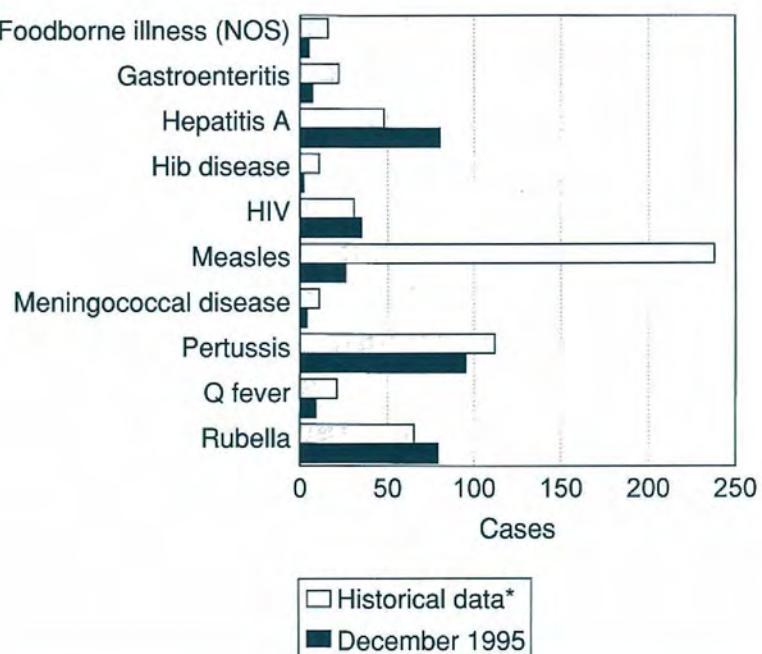
IDAC met for the first time in 1996 on January 31. One of its main roles is to review the Infectious Diseases Notification Schedule and to advise the Department on proposed changes. When the current schedule was introduced in 1991, the main criterion for inclusion of a condition on the list of notifiable diseases was that notification should lead to immediate public health action. For conditions such as meningococcal meningitis, measles and pertussis this has led to consistent action guided by clear protocols. For conditions such as syphilis and hepatitis C the difficulty in distinguishing prevalent and incident cases has made response more difficult. A further consequence of the more limited NSW schedule introduced in 1991 is that data for conditions such as chlamydia and campylobacter are no longer available to complement that of other States.

IDAC has recommended that the criteria for inclusion on the Notification Schedule be reviewed and that the

Continued on page 15 ►

FIGURE 7

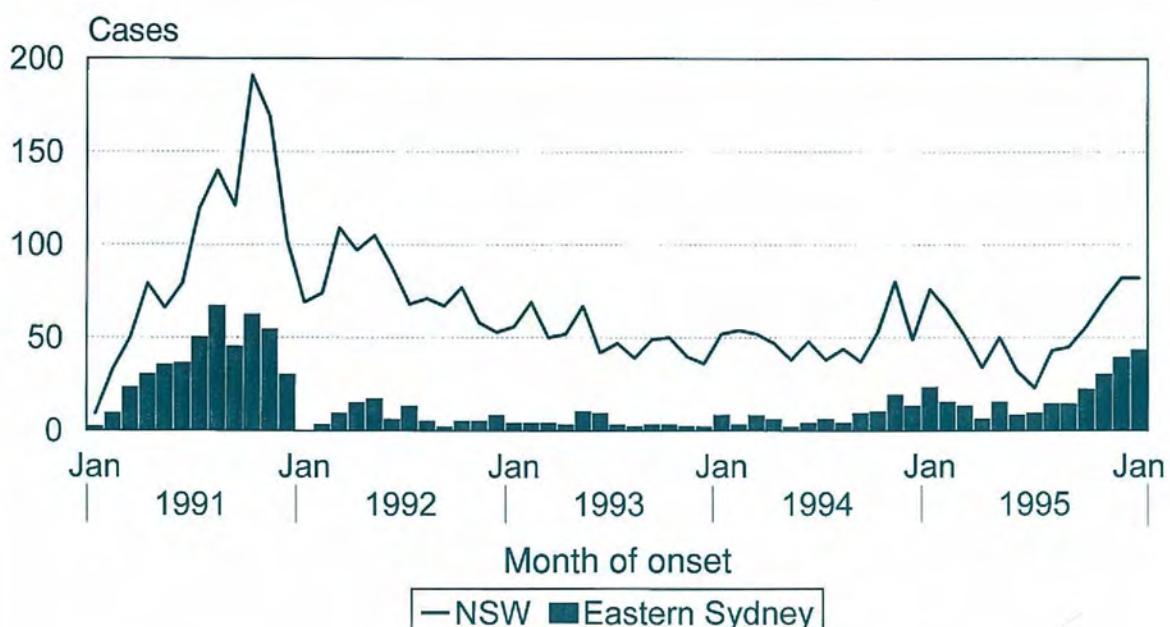
SELECTED INFECTIOUS DISEASES: NSW DECEMBER NOTIFICATIONS, 1995 COMPARED WITH HISTORICAL DATA



*Historical data: the average number of notifications diagnosed in the same month in the previous three years. Source: IDSS

FIGURE 8

HEPATITIS A NOTIFICATIONS FOR NSW 1991-1996, BY DATE OF ONSET



Source: IDSS

Infectious diseases

► Continued from page 13

Infectious Diseases Manual used by Public Health Units clearly distinguish between conditions for which an immediate response is required and those being notified to provide surveillance data to inform longer-term public health planning. The committee also recommended that the revision of the NSW schedule be carried out in consultation with the other States and Territories in support of a National Communicable Disease Surveillance Strategy.

Feedback by readers of the *Public Health Bulletin* to IDAC's recommendations would be welcome. Comments, criticisms and suggestions should be relayed to John Rooney, Specialist Medical Adviser, AIDS/Infectious Diseases Branch (e-mail jrooe@doh.health.nsw.gov.au, or fax (02) 391 9189).

TABLE 3

SUMMARY OF NSW INFECTIOUS DISEASE NOTIFICATIONS
DECEMBER 1995

Condition	Number of cases notified			
	Period		Cumulative	
	Dec 1994	Dec 1995	Dec 1994	Dec 1995
Adverse reaction	6	2	43	34
AIDS	46	20	598	413
Arboviral infection	9	15	384	542
Brucellosis	—	—	4	2
Cholera	—	—	—	1
Diphtheria	—	—	—	—
Foodborne illness (NOS)	14	5	232	389
Gastroenteritis (instit.)	35	7	310	1,374
Gonorrhoea	30	37	364	431
H influenzae epiglottitis	—	—	21	6
H influenzae B – meningitis	1	—	17	11
H influenzae B – septicaemia	1	1	12	9
H influenzae infection (NOS)	2	1	11	4
Hepatitis A	49	80	598	623
Hepatitis B	383	298	4,726	4,941
Hepatitis C	650	624	9,412	8,334
Hepatitis D	1	2	20	21
Hepatitis, acute viral (NOS)	—	—	2	2
HIV infection	20	35	428	452
Hydatid disease	1	2	20	18
Legionnaires' disease	4	7	61	74
Leprosy	—	—	3	2
Leptospirosis	1	1	14	6
Listeriosis	2	2	9	13
Malaria	14	4	187	96
Measles	266	26	1,505	599
Meningococcal meningitis	6	4	81	73
Meningococcal septicaemia	3	—	41	23
Meningococcal infection (NOS)	2	—	21	18
Mumps	1	1	11	14
Mycobacterial tuberculosis	28	19	415	427
Mycobacterial – atypical	45	3	523	395
Mycobacterial infection (NOS)	2	8	39	88
Pertussis	81	95	1,423	1,358
Plague	—	—	—	—
Poliomylitis	—	—	—	—
Q fever	25	9	268	202
Rubella	27	79	234	1,103
Salmonella infection (NOS)	107	93	1,086	1,275
Syphilis	64	40	1,072	902
Tetanus	—	—	4	—
Typhoid and paratyphoid	5	5	36	39
Typhus	—	—	—	—
Viral haemorrhagic fevers	—	—	—	—
Yellow fever	—	—	—	—

TABLE 4
**INFECTIOUS DISEASE CUMULATIVE NOTIFICATIONS FOR NSW, 1995
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996**

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	U/K	Total
AIDS	6	97	1	152	14	6	32	2	37	—	23	—	15	9	4	15	—	413
Arboviral infection	8	5	1	8	14	25	215	47	6	156	5	17	3	2	26	4	—	542
Brucellosis	1	—	—	—	—	—	—	—	—	—	—	1	—	—	—	—	—	2
Cholera	—	—	—	1	—	—	—	—	—	—	—	—	—	—	—	—	—	1
Gonorrhoea infection	5	71	14	179	7	17	19	12	20	8	23	1	17	4	23	11	—	431
Hepatitis B – acute viral	1	7	1	16	—	—	7	2	—	1	3	1	4	—	16	6	—	65
Hepatitis B – chronic/cARRIER	20	1	20	301	—	—	12	10	4	—	23	4	—	10	9	117	—	531
Hepatitis B – unspecified	32	539	12	58	102	101	61	17	547	34	681	16	1,439	45	13	648	—	4,345
Hepatitis C – acute viral	1	—	1	7	—	—	—	—	—	1	—	—	—	2	48	3	—	63
Hepatitis C – unspecified	211	918	340	1,186	505	494	842	237	590	276	530	237	930	195	38	739	—	8,271
Hepatitis D – unspecified	—	—	—	2	2	—	5	1	1	—	3	1	3	—	—	3	—	21
Hepatitis, acute viral (NOS)	—	—	—	1	—	—	—	—	—	—	—	—	—	—	1	—	—	2
HIV infection	12	81	3	177	18	11	7	2	26	1	12	5	30	7	1	20	39	452
Hydatid disease	—	1	1	1	1	1	1	—	1	2	3	2	3	—	—	1	—	18
Legionnaires' disease	2	3	—	6	15	6	1	3	9	—	2	—	3	3	2	19	—	74
Leprosy	—	1	—	—	—	—	—	—	1	—	—	—	—	—	—	—	—	2
Leptospirosis	—	—	—	—	—	—	—	—	—	—	—	—	—	—	—	—	—	6
Malaria	4	7	—	9	10	4	9	1	24	3	3	3	4	4	—	11	—	96
Meningococcal infection (NOS)	1	—	—	3	1	—	3	—	—	—	4	2	2	—	2	—	—	18
Meningococcal meningitis	8	1	5	5	11	11	4	4	8	4	3	—	5	1	—	3	—	73
Meningococcal septicaemia	—	4	—	—	5	1	2	1	2	1	1	1	2	—	—	—	—	23
Mycobacterial atypical	13	56	4	80	24	6	19	9	46	4	38	3	47	13	6	26	—	395
Mycobacterial infection (NOS)	7	14	—	3	2	—	4	—	11	1	17	—	19	2	—	8	—	88
Mycobacterial tuberculosis	6	42	1	27	8	6	5	3	49	2	54	4	113	6	5	96	—	427
Q fever	—	1	12	—	15	6	49	32	—	28	—	4	1	—	53	1	—	202
Syphilis infection	6	131	11	179	19	18	62	47	41	8	60	6	129	17	108	60	—	902
Vibrio infection (non cholera)	—	—	—	1	1	—	—	—	—	—	—	—	—	—	—	—	—	2

TABLE 5
**VACCINE PREVENTABLE AND RELATED CONDITIONS, CUMULATIVE NOTIFICATIONS FOR NSW, 1995
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996**

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	Total
Adverse event after immunisation	—	1	—	1	1	—	7	2	—	6	4	4	—	4	—	4	34
H. influenzae epiglottitis	—	—	1	1	—	—	1	—	—	2	—	—	—	—	1	—	6
H. influenzae infection (NOS)	1	—	—	—	1	—	1	—	—	—	—	—	—	—	1	—	4
H. influenzae meningitis	—	1	—	—	—	—	4	1	—	—	—	—	—	1	1	—	3
H. influenzae septicaemia	—	—	2	—	1	—	1	—	1	—	1	1	—	1	—	1	9
Measles	14	29	14	61	65	78	52	47	16	8	48	18	46	40	7	56	599
Mumps	—	—	—	2	—	2	3	—	2	1	2	—	—	—	—	2	14
Pertussis	39	26	23	35	77	100	342	18	114	46	70	123	88	119	12	126	1,358
Rubella	33	123	114	29	104	63	130	22	55	8	60	4	20	58	59	221	1,103

TABLE 6
**FOODBORNE INFECTIOUS DISEASE CUMULATIVE NOTIFICATIONS FOR NSW, 1995
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996**

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	Total
Foodborne illness (NOS)	16	9	3	3	188	1	3	3	4	—	1	8	101	—	26	23	389
Gastroenteritis (instit)	40	155	61	77	213	—	47	—	132	15	141	36	68	206	2	181	1,374
Hepatitis A – acute viral	15	93	39	207	26	10	32	4	53	2	43	11	44	5	3	38	625
Listeriosis	—	1	1	3	—	—	—	1	1	1	2	1	1	—	—	1	13
Salmonella (NOS)	29	69	32	94	93	63	132	77	139	43	127	28	110	70	44	125	1,275
Typhoid & paratyphoid	—	3	—	8	—	3	—	4	1	7	—	7	1	—	5	—	39

Abbreviations used in this Bulletin:

CSA Central Sydney Health Area, SSA Southern Sydney Health Area, ESA Eastern Sydney Health Area, SWS South Western Sydney Health Area, WSA Western Sydney Health Area, WEN Wentworth Health Area, NSA Northern Sydney Health Area, CCA Central Coast Health Area, ILL Illawarra Health Area, HUN Hunter Health Area, NC North Coast Public Health Unit, ND Northern District Public Health Unit, WN Western New South Wales Public Health Unit, CW Central West Public Health Unit, SW South West Public Health Unit, SE South East Public Health Unit, OTH Interstate/Overseas, U/K Unknown, NOS Not Otherwise Stated.

Please note that the data contained in this Bulletin are provisional and subject to change because of late reports or changes in case classification. Data are tabulated where possible by area of residence and by the disease onset date and not simply the date of notification or receipt of such notification.

TABLE 7**SUMMARY OF NSW INFECTIOUS DISEASE NOTIFICATIONS
JANUARY 1996**

Condition	Number of cases notified	
	Period	
	Jan 1995	Jan 1996
Adverse reaction	3	5
AIDS	45	17
Arboviral infection	21	22
Brucellosis	—	—
Cholera	—	—
Diphtheria	—	—
Foodborne illness (NOS)	17	12
Gastroenteritis (instit.)	2	—
Gonorrhoea	34	33
H influenzae epiglottitis	—	—
H influenzae B – meningitis	2	—
H influenzae B – septicaemia	—	—
H influenzae infection (NOS)	—	1
Hepatitis A	76	82
Hepatitis B	411	171
Hepatitis C	789	325
Hepatitis D	2	—
Hepatitis, acute viral (NOS)	—	3
HIV infection	50	40
Hydatid disease	—	1
Legionnaires' disease	16	2
Leprosy	1	—
Leptospirosis	1	2
Listeriosis	—	2
Malaria	22	18
Measles	98	18
Meningococcal meningitis	2	4
Meningococcal septicaemia	1	—
Meningococcal infection (NOS)	3	1
Mumps	2	3
Mycobacterial tuberculosis	50	6
Mycobacterial – atypical	47	1
Mycobacterial infection (NOS)	6	7
Pertussis	84	45
Plague	—	—
Poliomyelitis	—	—
Q fever	21	9
Rubella	34	21
Salmonella infection (NOS)	150	67
Syphilis	91	29
Tetanus	—	—
Typhoid and paratyphoid	6	6
Typhus	—	—
Viral haemorrhagic fevers	—	—
Yellow fever	—	—

TABLE 8**INFECTIOUS DISEASE NOTIFICATIONS FOR NSW, 1996
BY SELECTED MONTH OF ONSET FOR NOTIFICATIONS
RECEIVED BY JANUARY 31, 1996**

Condition	Oct	Nov	Dec	Jan	Total
Adverse event after immunisation	3	1	2	5	11
AIDS	41	28	20	17	106
Arboviral infection	7	22	15	22	66
Foodborne illness (NOS)	61	9	5	12	87
Gastroenteritis (instit.)	120	74	7	—	201
Gonorrhoea infection	37	37	37	33	144
H. influenzae infection (NOS)	—	1	1	1	3
H. influenzae meningitis	1	2	—	—	3
H. influenzae septicaemia	—	2	1	—	3
Hepatitis A – acute viral	56	70	80	82	288
Hepatitis B – acute viral	2	5	11	4	22
Hepatitis B – chronic/cARRIER	46	38	30	18	132
Hepatitis B – unspecified	396	363	257	149	1,165
Hepatitis C – acute viral	1	4	5	—	10
Hepatitis C – unspecified	678	666	619	325	2,288
Hepatitis D – unspecified	3	1	2	—	6
Hepatitis, acute viral (NOS)	1	—	—	3	4
HIV infection	29	37	35	40	131
Hydatid disease	2	2	2	1	7
Legionnaires' disease	3	3	7	2	15
Leptospirosis	0	1	1	2	4
Listeriosis	1	1	2	2	6
Malaria	2	4	3	18	27
Measles	52	36	26	18	132
Meningococcal infection (NOS)	3	—	—	1	4
Meningococcal meningitis	5	8	4	4	21
Meningococcal septicaemia	—	1	—	—	1
Mumps	—	4	1	3	8
Mycobacterial atypical	26	10	3	1	40
Mycobacterial infection (NOS)	9	13	8	7	37
Mycobacterial tuberculosis	39	22	19	6	86
Pertussis	146	122	95	45	408
Q fever	14	16	9	9	48
Rubella	199	157	79	21	456
Salmonella (NOS)	100	130	93	67	390
Syphilis infection	85	69	40	29	223
Typhoid and paratyphoid	2	1	5	6	14
Vibrio infection (non cholera)	—	—	1	—	1

TABLE 9

INFECTIOUS DISEASE CUMULATIVE NOTIFICATIONS FOR NSW, 1996
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	U/K	Total	
AIDS	1	-	-	2	1	-	1	-	11	-	-	-	1	-	-	-	-	17	
Arboviral infection	1	-	-	15	-	2	13	-	1	-	1	-	1	6	1	-	-	22	
Gonorrhoea infection	-	7	-	15	-	-	2	-	-	-	1	-	-	1	6	1	-	33	
Hepatitis B - acute viral	-	-	-	3	-	-	-	-	-	-	-	-	-	-	-	-	-	4	
Hepatitis B - chronic/cARRIER	-	-	-	8	-	-	1	-	-	-	-	-	-	-	1	3	-	18	
Hepatitis B - unspecified	5	-	-	3	-	-	3	-	27	-	16	-	26	4	-	24	-	149	
Hepatitis C - unspecified	1	31	-	14	3	-	3	44	1	39	6	12	17	17	25	2	34	-	325
Hepatitis, acute viral (NOS)	6	44	2	52	21	3	44	-	-	-	-	-	-	-	-	-	-	3	
HIV infection	-	-	1	1	-	-	1	-	-	-	-	-	-	-	-	1	29	40	
Hydatid disease	-	-	-	3	-	-	-	-	-	-	2	-	3	1	-	1	-	1	
Legionnaires' disease	-	-	-	-	1	-	-	-	-	-	-	-	1	1	-	-	-	2	
Leptospirosis	-	-	1	-	-	-	1	-	-	-	-	-	-	-	-	-	-	2	
Malaria	-	-	2	3	1	-	-	6	1	1	-	-	-	1	2	-	-	18	
Meningococcal infection (NOS)	-	-	-	-	-	-	1	-	-	-	-	-	-	-	-	-	-	1	
Meningococcal meningitis	-	-	-	-	1	1	-	-	-	-	1	-	-	-	1	-	-	4	
Mycobacterial atypical	-	-	-	-	-	-	-	1	-	-	-	-	-	-	-	-	-	1	
Mycobacterial infection (NOS)	2	-	-	-	2	-	1	-	-	-	-	-	-	-	-	2	-	7	
Mycobacterial tuberculosis	-	-	-	1	-	-	-	2	-	1	-	-	-	-	-	2	-	6	
Q fever	-	-	2	-	-	1	-	-	-	-	-	-	-	-	6	-	-	9	
Syphilis infection	-	4	3	4	1	1	-	6	2	1	-	-	-	-	2	4	-	29	

TABLE 10

VACCINE PREVENTABLE AND RELATED CONDITIONS, CUMULATIVE NOTIFICATIONS FOR NSW, 1996
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	Total
Adverse event after immunisation	-	-	2	-	-	-	-	-	-	2	-	-	-	1	-	-	5
H. influenzae infection (NOS)	-	-	1	-	-	-	-	-	-	-	-	-	-	-	-	-	1
Measles	-	-	-	-	-	4	-	-	-	1	1	3	-	1	1	7	18
Mumps	-	-	-	-	-	-	-	-	3	-	-	-	-	-	-	-	3
Pertussis	-	-	-	2	8	-	9	-	4	3	-	12	-	1	3	3	45
Rubella	-	7	-	-	-	1	-	-	-	1	-	-	-	1	-	11	21

TABLE 11

FOODBORNE INFECTIOUS DISEASE CUMULATIVE NOTIFICATIONS FOR NSW, 1996
BY PUBLIC HEALTH UNIT, RECEIVED BY JANUARY 31, 1996

Condition	CCA	CSA	CW	ESA	HUN	ILL	NC	ND	NSA	SE	SSA	SW	SWS	WEN	WN	WSA	Total
Foodborne illness (NOS)	-	-	-	-	-	-	-	-	-	-	-	2	7	-	3	-	12
Hepatitis A - acute viral	-	15	-	43	4	1	2	-	6	1	2	-	-	-	-	8	82
Listeriosis	-	-	-	-	-	-	-	-	-	1	-	-	1	-	-	-	2
Salmonella (NOS)	3	4	1	1	10	1	10	-	10	1	4	5	7	1	7	2	67
Typhoid and paratyphoid	-	2	-	-	1	-	-	-	-	-	-	2	-	-	1	6	